

$$X \sim U([-1, 1])$$

$$f(x) = \begin{cases} \frac{1}{2} & -1 < x < 1 \\ 0 & \text{else.} \end{cases}$$

$$X \sim U([a, b]) \quad f(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{else.} \end{cases}$$

$Y$  ind. de  $X$   $P(Y=1) = P(Y=-1) = \frac{1}{2}$

$U = XY$  he la stessa legge di  $X$

$$P(U \leq t) = P(X \leq t)$$

$$\boxed{Z = X + U} \quad \text{non } \bar{e} \text{ una v.a. continua}$$

$$\boxed{P(Z=0)} = P(X+U=0) =$$

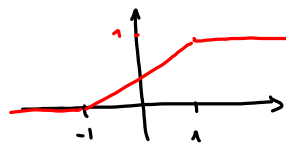
$$= P(\underbrace{X(1+Y)}=0) = P(1+Y=0)$$

$$= P(Y = -1) = \frac{1}{2}$$

(c) Calc. la f.d.a. di  $Z = X(1+Y)$

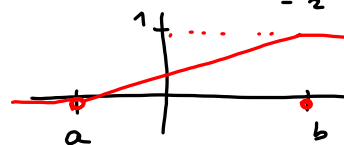
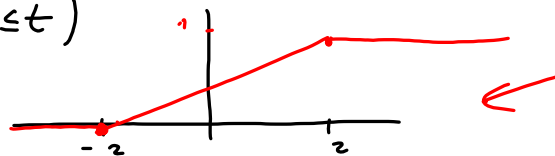
$$\begin{aligned}
 P(Z \leq t) &= P(X(1+Y) \leq t) = \\
 &= P(X(1+Y) \leq t, Y=1) + \\
 &\quad + P(X(1+Y) \leq t, Y=-1) = \\
 &= P(\underbrace{2X \leq t}, \underbrace{Y=1}) + P(0 \leq t, Y=-1) \\
 &= P(2X \leq t) P(Y=1) + P(0 \leq t) P(Y=-1) \\
 &= \frac{1}{2} P(X \leq \frac{t}{2}) + \frac{1}{2} P(0 \leq t)
 \end{aligned}$$

$$P(X \leq u) = \begin{cases} 0 & u < -1 \\ \frac{u+1}{2} & -1 \leq u \leq 1 \\ 1 & u > 1 \end{cases}$$



$$P(X \leq \frac{t}{2}) = \begin{cases} 0 & t < -2 \\ \frac{\frac{t}{2} + 1}{2} = \frac{t+2}{4} & -2 \leq t \leq 2 \\ 1 & t > 2 \end{cases}$$

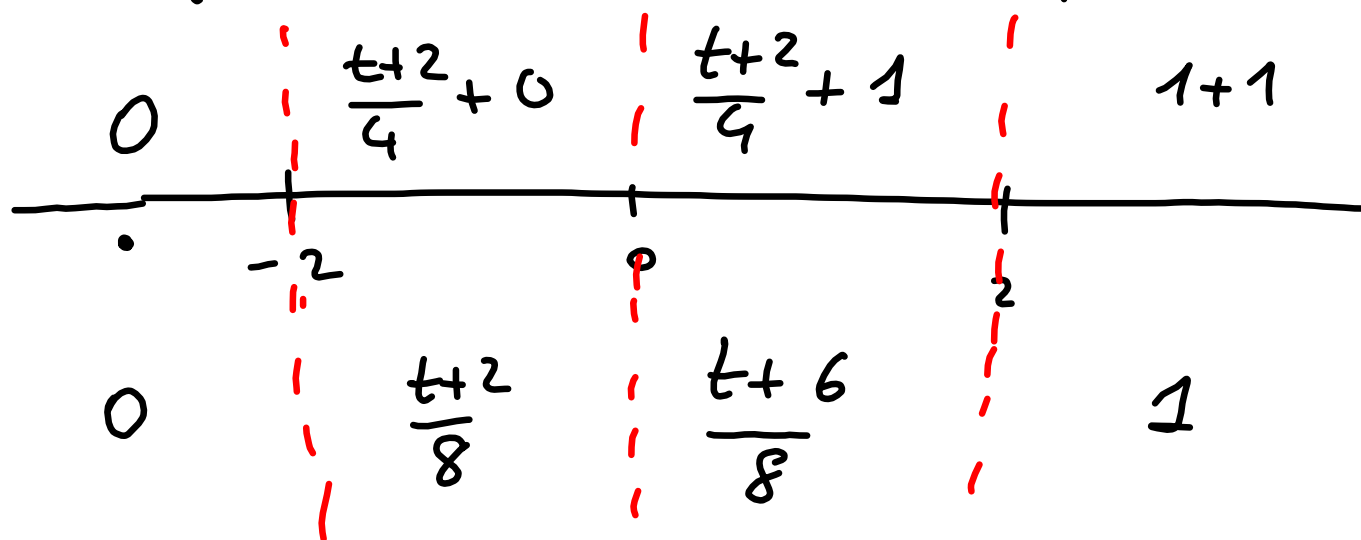
$$P(2X \leq t)$$



$2X \sim U(-2, 2)$

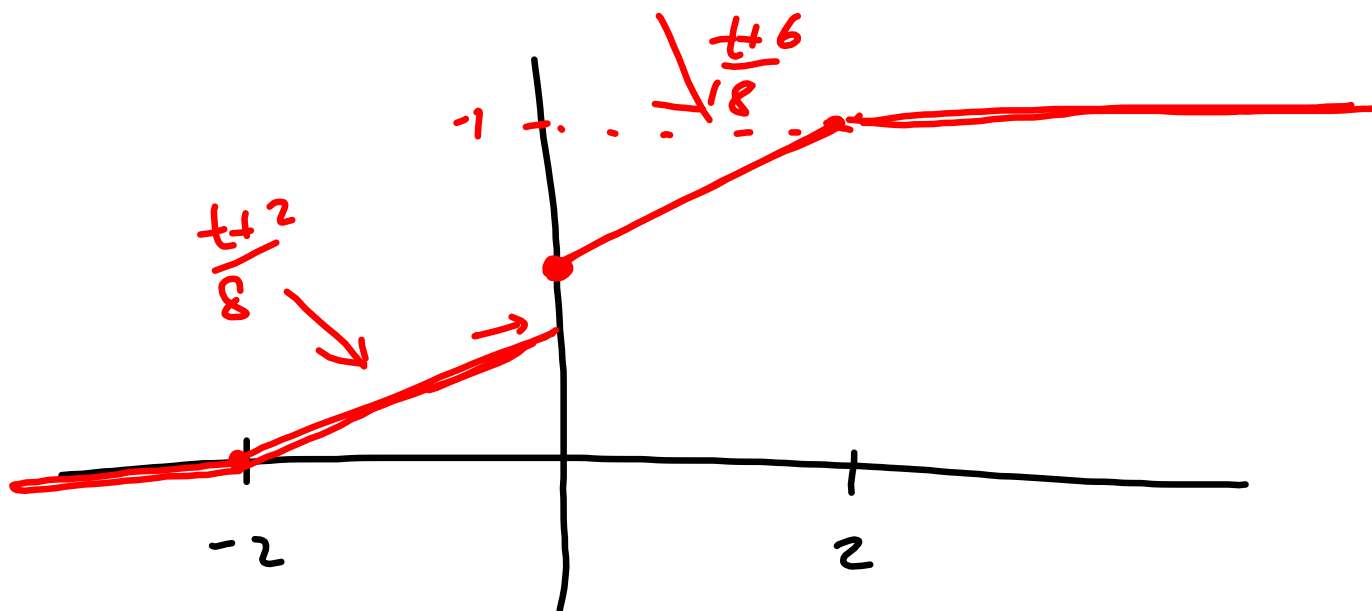
$$\frac{1}{2} P(X \leq \frac{t}{2}) + \frac{1}{2} P(0 \leq t) =$$

$$= \frac{1}{2} \begin{cases} 0 & t < -2 \\ \frac{t+2}{4} & -2 \leq t \leq 2 \\ 1 & t > 2 \end{cases} + \frac{1}{2} \begin{cases} 0 & t < 0 \\ 1 & t \geq 0 \end{cases}$$



$$0 \quad \frac{t+2}{8} \quad \frac{t+6}{8} \quad 1$$

fare il grafico  $P(Z=0)$



$$P(z=0) = \frac{6}{8} - \frac{2}{8} = \frac{1}{2}$$

$L \rightarrow \frac{1}{3}$   
 $B \rightarrow \frac{2}{3}$

- Se il segnale è L, il tempo in attesa su tempo aleatorio  $\sim \frac{2}{3}(2)$
- Se il segnale è B, il tempo  $\sim \frac{2}{3}(1)$

$Z =$  tempo in attesa dal segnale per giungere a destinazione

(a) Calcolare  $P(Z \leq t | B)$   $P(Z \leq t | L)$   
 dove  $B = \{ \text{il segnale è B} \}$   $L = \{ \text{il segnale è L} \}$

(b) Calcolare la f.d.n. di  $Z \in \mathbb{R}^+$

(a)  $P(Z \leq t | B) = \begin{cases} 0 & t < 0 \\ 1 - e^{-t} & t \geq 0 \end{cases}$

$P(Z \leq t | L) = \begin{cases} 0 & t < 0 \\ 1 - e^{-2t} & t \geq 0 \end{cases}$

---

$P(Z \leq t) = ?$

$P(A) = P(A|B)P(B) + P(A|L)P(L)$

$= \frac{2}{3} \begin{cases} 0 & t < 0 \\ 1 - e^{-t} & t \geq 0 \end{cases} + \frac{1}{3} \begin{cases} 0 & t < 0 \\ 1 - e^{-2t} & t \geq 0 \end{cases} =$

$= \begin{cases} 0 & t < 0 \\ \frac{2}{3}(1 - e^{-t}) + \frac{1}{3}(1 - e^{-2t}) & t \geq 0 \end{cases} = 1 - \frac{2}{3}e^{-t} - \frac{1}{3}e^{-2t}$

$F(t) = \int_{-\infty}^t f(x) dx = \int_{-\infty}^a f(x) dx + \int_a^t f(x) dx$

$F(t) = c + \int_a^t f(x) dx$

$F'(t) = f(t)$

---

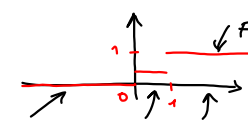
Teor. Fond. del calcolo  
 Sia  $f$  una funzione continua su  $[a, b]$ .

$F(t) = \int_a^t f(x) dx$

allora  $F$  è derivabile in  $(a, b)$  e

$F'(t) = f(t) \quad \forall t \in (a, b)$

$F' = f$   
 $F(t) = \int_{-\infty}^t f(x) dx \quad \forall t$



$f(x) = 0 \quad \forall x$

---

$F(t) = P(Z \leq t) = \begin{cases} 0 & t < 0 \\ 1 - \frac{2}{3}e^{-t} - \frac{1}{3}e^{-2t} & t \geq 0 \end{cases}$

$f(t) = \begin{cases} 0 & t \leq 0 \\ \frac{2}{3}e^{-t} + \frac{2}{3}e^{-2t} & t > 0 \end{cases}$

$\int_{-\infty}^t f(x) dx \quad \forall t$

$E[Z] = \int_{-\infty}^{+\infty} t f(t) dt =$

$= \int_0^{+\infty} t \left( \frac{2}{3}e^{-t} + \frac{2}{3}e^{-2t} \right) dt =$

$= \frac{2}{3} \int_0^{+\infty} t e^{-t} dt + \frac{1}{3} \int_0^{+\infty} 2t e^{-2t} dt$

$= \frac{2}{3} \cdot 1 + \frac{1}{3} \cdot \left( \frac{1}{2} \right) = \frac{2}{3} + \frac{1}{6}$

Let  $X \sim \mathcal{E}(\lambda)$

$E[X] = \int_0^{+\infty} t \lambda e^{-\lambda t} dt = \frac{1}{\lambda}$

nov 21-15:06

Soit  $X \sim \mathcal{E}(\lambda)$

- (a) Calc. la f.d.r. de  $Y = \frac{1}{X}$   
 (b) Calc. la densité de  $Y$   
 (c) calc.  $E[Y]$ .

(a)  $P(Y \leq t) = P\left(\frac{1}{X} \leq t\right) = \underline{P(1 \leq Xt)}$

$$P(X \leq t) = \begin{cases} 0 & t \leq 0 \\ 1 - e^{-\lambda t} & t > 0 \end{cases} \quad P(X \leq 0)$$

$$X \sim f(x) = \begin{cases} \lambda e^{-\lambda x} & x > 0 \\ 0 & x \leq 0 \end{cases}$$

$$P\left(\frac{1}{X} \leq t\right) = P\left(\frac{1}{X} \leq t, X > 0\right)$$

$$+ P\left(\frac{1}{X} \leq t, X < 0\right) =$$

$$= P(1 \leq Xt, X > 0) + P(1 \geq Xt, X < 0)$$

---


$$\underline{P(Y \leq t)} = P\left(\frac{1}{X} \leq t\right) = P(1 \leq Xt)$$

$$= P\left(\frac{1}{t} \leq X\right) \quad t > 0$$

Soit  $t \leq 0$   
 $P\left(\frac{1}{X} \leq t\right)$



$$X \sim \mathcal{E}(\lambda) \quad Y = \frac{1}{X}$$

Dato che  $P(X > 0) = 1$ , anche  $P(\frac{1}{X} > 0)$

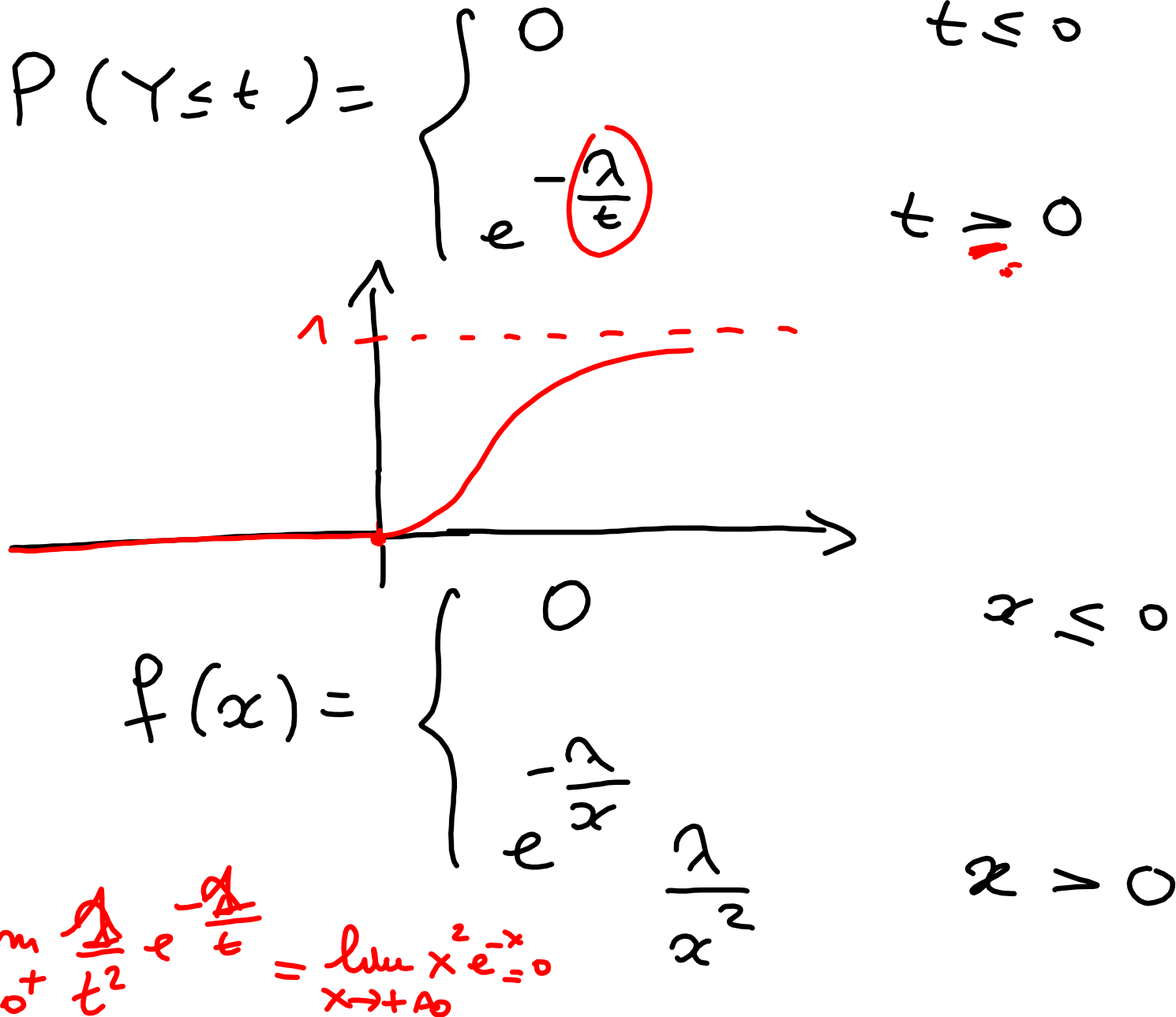
Diunque  $P(\frac{1}{X} \leq t) = 0$  se  $t \leq 0$

→ Se  $t > 0$

$$P(\frac{1}{X} \leq t) = P(1 \leq tX) = P(X \geq \frac{1}{t}) =$$

$$= 1 - P(X < \frac{1}{t}) = 1 - P(X \leq \frac{1}{t}) =$$

$$= 1 - \begin{cases} 0 & \frac{1}{t} \leq 0 \\ 1 - e^{-\lambda \frac{1}{t}} & \frac{1}{t} > 0 \end{cases} = e^{-\frac{\lambda}{t}}$$



$$\begin{aligned}
 E[Y] &= \int_{-\infty}^{+\infty} y f(y) dy = \\
 &= \int_0^{+\infty} y \frac{\lambda}{y^2} e^{-\frac{\lambda}{y}} dy = \\
 &= \int_0^{+\infty} \frac{\lambda}{y} e^{-\frac{\lambda}{y}} dy \\
 &= - \int_0^{+\infty} \lambda u e^{-\lambda u} \frac{1}{u^2} du = \\
 &= \int_0^{+\infty} \frac{\lambda}{u} e^{-\lambda u} du = +\infty
 \end{aligned}$$

$$\begin{aligned}
 \frac{1}{y} &= u \\
 y &= \frac{1}{u} \\
 dy &= -\frac{1}{u^2} du
 \end{aligned}$$

$$\Gamma(x) = \int_0^{+\infty} u^{x-1} e^{-\lambda u} du \quad x > 0$$

$$X \sim \mathcal{P}(\lambda) \quad Y = \frac{1}{X} = \varphi(X)$$

$$E[Y] = E[\varphi(X)] =$$

$$= \int_{-\infty}^{+\infty} \varphi(x) f(x) dx = \int_0^{+\infty} \frac{1}{x} \lambda e^{-\lambda x} dx$$

dove  $f$  è la densità di  $X$

$$\text{Let } X \sim U([a, b]), \text{ then}$$
$$\text{the r.v. } Y = mX + n \sim U([ma+n, mb+n])$$
$$2x \quad U([2a, 2b])$$